Dete	ection Algo	rithms fo	r								
BIOSI	irveillance:	A tutori	al								
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	Tutorial slides by Andrew Moore										
Note to other teachers and users of these si found this source material useful in giving y these sildes verbatim, or to modify them to originals are available. If you make use of a your own lecture, please include this messa repository of Andrew's tutorials: http://www Comments and corrections gratefully receive Converted to 2002. 2002, Andrew M	lides. Andrew would be delighted if you our own lectures. Feel free to use fit your own needs. PowerPoint significant portion of these sildes in ge, or the following link to the source <u>tes crau edu/-awm/tutorials</u> . d.	RODS: <u>http://www.h</u> Auton Lab: <u>http://</u>	ealth.pitt.edu/rods www.autonlab.org								

	nas	Tried	Tried	Under development	Multivariate	Spatial
	Pitt/CMU	but little	and		signal	?
	tried it?	used	used		tracking?	
Time-weighted averaging	Yes	Yes				
Serfling	Yes		Yes			
ARIMA	Yes	Yes				
SARIMA + External Factors	Yes		Yes			
Jnivariate HMM	Yes		Yes			
Kalman Filter	Yes	Yes				
Recursive Least Squares	Yes		Yes			
Support Vector Machine	Yes	Yes				
Neural Nets	Yes	Yes				
Randomization	Yes		Yes	Yes		
Spatial Scan Statistics	Yes			(w/ Howard Burkom)	Yes	Yes
Bayesian Networks	Yes			Yes	Yes	
Contingency Tables	Yes		Yes			
Scalar Outlier (SQC)	Yes	Yes				
Multivariate Anomalies	Yes		Yes		Yes	
Change-point statistics	Yes			Yes		
DR Tests	Yes		Yes		Yes	
WSARE (Recent patterns)	Yes		Yes	Yes	Yes	Yes
PANDA (Causal Model)	Yes			Yes	Yes	Yes
LUMOD (space/Time HMM)				Yes	Yes	Yes















































Performance Market of the state of the stat	Performance Andress Andres Andress Andress	Performance standard control chart using yesterday Moving Average 7 Moving Average 56 0.54 2.72 0.44 0.39 0.36 0.36 0.47 0.22 4.13 0.39 0.36 0.347 0.22 4.13 0.33 3.79 0.58 2.79 0.51 3.31 Moving Average 56 0.54 2.72 0.44 3.54	Performance standard control chart using yesterday Moving Average 7 Moving Average 56 0.54 2.72 0.44 0.39 0.36 0.47 0.22 4.13 0.39 0.36 0.47 0.22 4.13 0.37 0.33 3.79 0.58 2.79 0.51 3.31 Moving Average 56 0.54 2.72 0.44 3.54	Performance	Allowing one False per TWO weeks	Alarm	Allowin	g one Fa (weeks.	alse Alar
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	58	2.79	0.51	3.31
Moving Average 56 0.4	54	2.72	0.44	3.54
hours_of_daylight 0.4	58	2.73	0.43	3.9
hours_of_daylight is_mon	0.7	2.25	0.57	3.12
hours_of_daylight is_mon is_tue 0.	72	1.83	0.57	3.16
hours_of_daylight is_mon is_sat 0.	77	2.11	0.59	3.26
CUSUM 0.4	45	2.03	0.15	3.55





that the date is a Sunday. It is also possible to add local terms, such as the mean count over the previous seven days as an additional feature to allow the method to adapt better to recent local events. These additions can be very helpful, and as we shall see at the end of the chapter, regression methods that include these extra terms perform better than moving average (which had been our favorite method) on our data.

10. SICKNESS AVAILABILITY

Another way to deal with day-of-week variations is to use the sickness availability method to smooth the time series by removing noise due to the day-of-week effect. This algorithm transforms the daily counts in the time series into a daily sickness value, which is defined as the number of people getting sick every day irrespective of whether they seek health care or not. The term availability refers to the probability that a patient will seek health care during a specific day of the week; hence, there are a total of seven values of availability, one for each day of the week. The availability of a day can be thought of as the fraction of a weeks-worth of visits that get assigned to the given day. The sickness availability method is based on the intuitive assumptions that the expected count is the product of the true amount of sickness and the current day's availability.

We can estimate the expected availability for a specific day of week (dow) using the average of the availabilities on that day for the past m weeks. We can calculate the expected availability A_{dow} as:

$$A_{dow} = \frac{\sum_{i=1}^{m} (C_{i(dow)} / \sum_{dow=0}^{6} C_{i(dow)})}{m}$$
(9)

In Eq. 9, A_{dow} refers to the expected availability for the day of week specified by the variable *dow*, which takes on seven different values ranging from 0 to 6. The term $C_{i(dow)}$ is the actual number of patients that visited the ED on the particular day of week *dow* during the *i*th week in the past. Since national holidays affect the number of patients visiting the ED, weeks containing holidays are ignored completely in the availability calculations. Finally, the parameter *m* controls the smoothness of the sickness curve.

Since sickness is defined as the total number of people in the city getting sick on a particular day, we can calculate it as follows:

$$S_{today} = \frac{C_{today}}{A_{today}} \tag{10}$$

The term S_{today} in Eq. 10 refers to the number of people who are sick on the current day while the term C_{today} refers to the number of people who visited the ED on the current day. A_{today} is the probability that a patient will visit the ED for the day of week specified by *today*. Figure 14.18 shows the availability estimated in the period leading up to the synthetic ramp outbreak: it shows a consistent picture that we usually see far more patients on Mondays than Sundays, and then the visits taper off during the rest of the week. Figure 14.19 shows both the original count and the estimated sickness (count/availability). Sickness is much more stable than the original count because day of week effects have been greatly reduced. It is now possible to run a time series algorithm on the sickness values. In this case, we chose to use moving average with a window of seven days. The resulting alarms show something that was not achieved in any





FIGURE 14.19 The black "count" line shows the raw data. The gray "sickness" line shows the sickness after the count has been divided by the corresponding availability value from Figure 14.18. Note how the sickness time series is now smoother than the original data and decorrelated with day of week. The alarm levels are derived from moving average applied to the sickness time series.

of the previous illustrations: a strong alarm resulting from the higher-than-expected counts for the Sunday.

We would like to emphasize that the sickness availability method only smoothes the time series. It is not a detection algorithm by itself. Instead, a detection algorithm, such as the control chart, moving average, or CUSUM algorithms should be used on the smoothed data.

11. FURTHER COMPARISON OF THE UNIVARIATE ALGORITHMS

We insert another copy of Table 14.1, with some of the above methods added for comparison. The newly evaluated algorithms follow:

- Regression using two features: the mean count over the past week and hours of daylight. This allows the algorithm to account for seasonal variation by putting a negative coefficient in front of hours of daylight.
- Regression using the additional feature *is_Monday*, which is set to 1 if today is Monday and 0 otherwise. This allows the algorithm to anticipate the Monday bump in physician visits and so be less prone to false positives.
- Regression using indicator variables for all days of the week except for Sunday (which would be redundant), and additionally, *hours_of_daylight* and *mean_count_over_previous_seven_days*.

 Using sickness/availability to compensate day-of-week effects, and then using the approach of comparing against yesterday. This method thus looks for jumps in the day-ofweek-adjusted counts.

For this data set, with its seasonal components and day-of-week components, we see that sickness availability (to cope with dayof-week effects) combined with moving average (to cope with seasonal trends) performs well. Some of the regression methods perform almost equally well. We should note that this does not mean these methods are best in general: individual properties of individual data sets mean different approaches can be stronger for different data sets. Our only general advice is that in our experience relatively simple methods usually work at least as well as complex approaches. A second important note is that the numbers in Table 14.2 cannot be used as an estimate of how quickly real outbreaks are expected to be detected: the numbers are a function of many things, including the simulated magnitude of the outbreaks and the simulated noise levels.

12. ADDITIONAL METHODS

A complete set of approaches would require a very thick book, such as Hamilton (1994). The purpose of this chapter has been a tour of a sufficient variety of approaches to introduce the reader to some of the issues faced when choosing or implementing a time-series-based method, without going into the













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using yesterday	0.14	3.83	0.1	4.7			
Moving Average 3	0.36	3.45	0.33	3.79			
Moving Average 7	0.58	2.79	0.51	3.31			
Moving Average 56	0.54	2.72	0.44	3.54			
hours_of_daylight	0.58	2.73	0.43	3.9			
hours_of_daylight is_mon	0.7	2.25	0.57	3.12			
hours_of_daylight is_mon is_tue	0.72	1.83	0.57	3.16			
hours_of_daylight is_mon is_sat	0.77	2.11	0.59	3.26			
CUSUM	0.45	2.03	0.15	3.55			
sa-mav-1	0.86	1.88	0.74	2.73			
sa-mav-7	0.87	1.28	0.83	1.87			
	0.86	1.27	0.82	1.62			

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Moving Average 7	0.58	2.79	0.51	3.31
Moving Average 56	0.54	2.72	0.44	3.54
hours_of_daylight	0.58	2.73	0.43	3.9
hours_of_daylight is_mon	0.7	2.25	0.57	3.12
hours_of_daylight is_mon is_tue	0.72	1.83	0.57	3.16
hours_of_daylight is_mon is_sat	0.77	2.11	0.59	3.26
CUSUM	0.45	2.03	0.15	3.55
sa-mav-1	0.86	1.88	0.74	2.73
sa-mav-7	0.87	1.28	0.83	1.87
sa-mav-14	0.86	1.27	0.82	1.62
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Moving Average 7	0.58	2.79	0.51	3.31
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hours_of_daylight	0.58	2.73	0.43	3.9
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hours_of_daylight is_mon is_tue	0.72	1.83	0.57	3.16
hours_of_daylight is_mon is_sat	0.77	2.11	0.59	3.26
CUSUM	0.45	2.03	0.15	3.55
sa-mav-1	0.86	1.88	0.74	2.73
sa-mav-7	0.87	1.28	0.83	1.87
sa-mav-14	0.86	1.27	0.82	1.62
sa-regress	0.73	1.76	0.67	2.21
Cough with denominator	0.78	2.15	0.59	2.41
Cough with MA	0.65	2.78	0.57	3.24
opyright © 2002, 2003, Andrew Moore	BIOSUI	veillance De	etection Aig	orithms: Silde































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Simple WSARE						
	Date	Cases				
	Thu 5/22/2000	C1, C2, C3, C4				
 Given 500 day's 	Fri 5/23/2000	C1, C2, C3, C4				
worth of ED cases at	:	: : C1, C2, C3, C4				
15 hospitals	:					
	Sat 12/9/2000					
	Sun 12/10/2000	C1, C2, C3, C4				
	:	:				
	Sat 12/16/2000	C1, C2, C3, C4				
	:	:				
	Sat 12/23/2000	C1, C2, C3, C4				
	:	:				
	:	:				
	Fri 9/14/2001	C1, C2, C3, C4				
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Simple	WSARE				
	Date	Cases			
	Thu 5/22/2000	C1, C2, C3, C4			
 Given 500 dav's 	Fri 5/23/2000	C1, C2, C3, C4			
worth of ED cases at	:	:			
	:	:			
15 hospitals	Sat 12/9/2000	C1, C2, C3, C4 C1, C2, C3, C4			
 For each day 	Sun 12/10/2000				
	:	:			
 Take today's cases 	Sat 12/16/2000	C1, C2, C3, C4			
 The cases one week ago 	:	:			
 The cases two weeks ago 	Sat 12/23/2000	C1, C2, C3, C4			
5	:	:			
	:	:			
	Fri 9/14/2001	C1, C2, C3, C4			
Copyright © 2002, 2003, Andrew Moore	Biosurveilland	e Detection Algorithms: Slide 88			

Example							
<pre>Sat 12-23-2001 (daynum 36882, dayindex 239) FISHER_PVALUE = 0.000051 35.8% (48/134) of today's cases have 30 <= age < 40 17.0% (45/265) of other cases have 30 <= age < 40</pre>							
Table 1: A sample 2x2 Contingency Table C_{today} C_{other} $Age_Decile = 3$ 4845 $Age_Decile \neq 3$ 86220							
Copyright © 2002, 2003, Andrew Moore Biosurveillance Detection Algorithms: Slide 92							

Example						
<pre>Sat 12-23-2001 (daynum 36882, dayindex 239) FISHER_PVALUE = 0.000051 RANDOMIZATION_PVALUE = 0.031 35.8% (48/134) of today's cases have 30 <= age < 40 17.0% (45/265) of other cases have 30 <= age < 40</pre>						
Table 1: A sample 2x2 Contingency Tab C_{today} C_{othe} $Age_Decile = 3$ 4845 $Age_Decile \neq 3$ 86220	le r					
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Checking twTable 2: 2x2 Contingency TableRecords from Today matching C_0 and C_1 Records from Today matching C_1 and differing on C_0 Table 3: 2x2 Contingency TableRecords from Today matching C_0 and C_1 Records from Today matching C_0 and differing on C_1	COMPONE e 1 for a two component rule Records from Other matching C_0 and C_1 Records from Other matching C_1 and differ- ing on C_0 e 2 for a two component rule Records from Other matching C_0 and C_1 Records from Other matching C_0 and differ- ing on C_1	ent •	rules Must pass both tests to be allowed to live.
--	--	----------	--

WSARE v2.0																
• Ir	Inputs: 1. Date/time-indexed biosurveillance-			2. Time Window Length				3. Which attributes to use?								
Outputs: 1. Here are the records that most surprise me				2	2. Here's why should take it					you						
Primary	Primary Date Time		Hospital ICD9 Prodrome Gender			iender Age H	Home		Work			Recent	Recent	(Many		
Кеу								Large Scale	Medium Scale	Fine Scale	Large Scale	Medium Scale	Fine Scale	Flu Levels	Weather	more)
h6r32	6/2/2	14:12	Down- town	781	Fever	М	20s	NE	15217	A5	NW	15213	B8	2%	70R	
t3q15	6/2/2	14:15	River- side	717	Respirat ory	М	60s	NE	15222	13	NE	15222	J3	2%	70R	
t5hh5	6/2/2	14:15	Smith- field	622	Respirat ory	F	80s	SE	15210	К9	SE	15210	К9	2%	70R	
:	:	:	:	:	:	:	:	:	:	:	:	:	:	:	:	:
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WSARE on recent Utah Data

Saturday June 1st in Utah:

The most surprising thing about recent records is:

Normally:

0.8% of records (50/6205) have time before 2pm and prodrome = Hemorrhagic But recently:

2.1% of records (19/907) have time before 2pm and prodrome = Hemorrhagic

Pvalue = 0.0484042

Which means that in a world where nothing changes we'd

expect to have a result this significant about once

every 20 times we ran the program

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	Reference	es				
1.	WSARE 3.0 : Bayesian Networ Detection	k based Anomaly Pattern				
	Wong, Moore, Cooper and Wag	gner [ICML/KDD 2003]				
2.	Fast Grid Based Computation of	of Spatial Scan Statistics				
	Neill and Moore [NIPS 2003]					
	These and other Biosurveillance algorithms papers and free software available from					
	http://www.autonlab.org/					
See also: http://www.health.pitt.edu/rods						
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