



Lecture 5

Subspace Transformations

Eigendecompositions, kernel PCA and CCA

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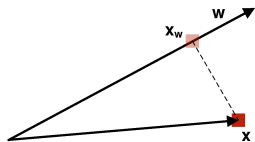
Advanced Topics in Machine Learning, 2012

Recall: Projections



- Projection of a point \mathbf{x} onto a direction \mathbf{w} is computed as:

$$\text{proj}_{\mathbf{w}}(\mathbf{x}) = \mathbf{w} \frac{\mathbf{w}^T \mathbf{x}}{\|\mathbf{w}\|^2}$$



- Directions in an RKHS expressed as linear combination of points:

$$\mathbf{w} = \sum_{i=1}^N \alpha_i \phi(\mathbf{x}_i)$$

- The norm of the projection onto \mathbf{w} thus can be expressed as

$$\|\text{proj}_{\mathbf{w}}(\mathbf{x})\| = \frac{\mathbf{w}^T \mathbf{x}}{\|\mathbf{w}\|} = \frac{\sum_{i=1}^N \alpha_i \kappa(\mathbf{x}_i, \mathbf{x})}{\sqrt{\sum_{i,j=1}^N \alpha_i \alpha_j \kappa(\mathbf{x}_i, \mathbf{x}_j)}} = \sum_{i=1}^N \beta_i \kappa(\mathbf{x}_i, \mathbf{x})$$

Thus, the *size* of the projection onto \mathbf{w} can be expressed as a linear combination of the kernel valuations with \mathbf{x}



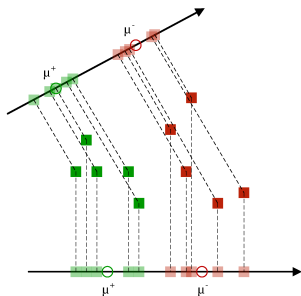
- In LDA, we chose a projection direction \mathbf{w} to maximize the cost function

$$J(\mathbf{w}) = \frac{\|\mu_{\mathbf{w}}^+ - \mu_{\mathbf{w}}^-\|^2}{(\sigma_{\mathbf{w}}^+)^2 + (\sigma_{\mathbf{w}}^-)^2} = \frac{\mathbf{w}^T \mathbf{S}_B \mathbf{w}}{\mathbf{w}^T (\mathbf{S}_W^+ + \mathbf{S}_W^-) \mathbf{w}}$$

where μ^+ & μ^- are the averages of the sets, σ^+ & σ^- are their standard deviations, \mathbf{S}_B is the **between scatter matrix** & \mathbf{S}_W^+ and \mathbf{S}_W^- are the **within scatter matrices**

- The optimal solution \mathbf{w}^* is given by the first eigenvector of the matrix

$$(\mathbf{S}_W^+ + \mathbf{S}_W^-)^{-1} \mathbf{S}_B$$





- When the projection direction is in feature space, $\mathbf{w}_\alpha = \sum_{i=1}^N \alpha_i \phi(\mathbf{x}_i)$
- From this, the LDA objective can be expressed as

$$\max_{\alpha} J(\alpha) = \frac{\alpha^\top \mathbf{M} \alpha}{\alpha^\top \mathbf{N} \alpha}$$

where

$$\mathbf{M} = (\mathbf{K}_+ - \mathbf{K}_-) \mathbf{1}_N \mathbf{1}_N^\top (\mathbf{K}_+ - \mathbf{K}_-)$$

$$\mathbf{N} = \mathbf{K}_+ \left(\mathbf{I}_{N^+} - \frac{1}{N^+} \mathbf{1}_{N^+} \mathbf{1}_{N^+}^\top \right) \mathbf{K}_+^\top + \mathbf{K}_- \left(\mathbf{I}_{N^-} - \frac{1}{N^-} \mathbf{1}_{N^-} \mathbf{1}_{N^-}^\top \right) \mathbf{K}_-^\top$$

- Solutions α^* to the above generalized eigenvalue problem (as discussed later) allow us to project data onto this discriminant direction as

$$\|\text{proj}_{\mathbf{w}}(\mathbf{x})\| = \sum_{i=1}^N \alpha_i^* \kappa(\mathbf{x}_i, \mathbf{x})$$



- **Objective:** find a subspace that captures an important aspect of the training data... we find K axes that span this subspace
- **General Problem:** we will solve problems

$$\max_{g(\mathbf{w})=1} f(\mathbf{w})$$

for projection direction \mathbf{w} ... iteratively solving these problems will yield a subspace defined by $\{\mathbf{w}_k\}_{k=1}^K$

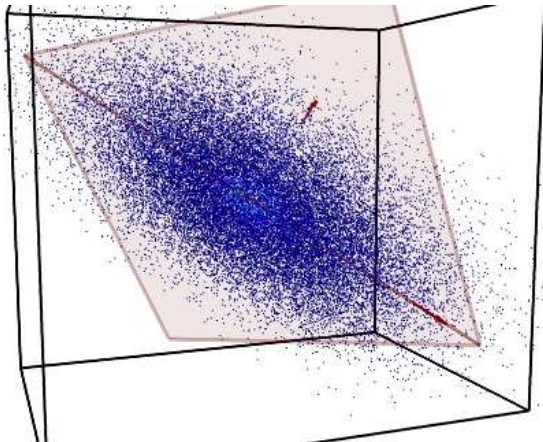
- **General Approach:** find a center $\boldsymbol{\mu}$ and a set of K orthonormal directions $\{\mathbf{w}_k\}_{k=1}^K$ used to project data into the subspace:

$$\tilde{\mathbf{x}} \leftarrow \left(\mathbf{w}_k^\top (\mathbf{x} - \boldsymbol{\mu}) \right)_{k=1}^K$$

- This is a K -dimensional representation of the data *regardless* of the original space's dimensionality—the coordinates in the space spanned by $\{\mathbf{w}_k\}_{k=1}^K$
- This projection will be centered at $\mathbf{0}$ (in feature space)



We want to find subspace that captures important aspects of our data





- LDA found 1 direction for discriminating between 2 classes
- In this lecture, we will see 3 subspace projection objectives / techniques:
 - Find directions that maximize variance in X (PCA)
 - Find directions that maximize covariance between X & Y (MCA)
 - Find directions that maximize correlation X & Y (CCA)
- These techniques extract underlying structure from the data allowing us to...
 - Capture fundamental structure of the data
 - Represent the data in low dimensions
- Each of these techniques can be kernelized to operate in a feature space yielding kernelized projections onto \mathbf{w} :

$$\|\text{proj}_{\mathbf{w}}(\phi(\mathbf{x}))\| = \mathbf{w}^T \phi(\mathbf{x}) = \sum_{i=1}^N \alpha_i \kappa(\mathbf{x}_i, \mathbf{x}) \quad (1)$$

where α is the vector of dual values defining \mathbf{w}

Part I

Principal Component Analysis

Motivation: Directions of Variance



- We want to find a direction \mathbf{w} that maximizes the data's variance
- Consider a random variable $\mathbf{x} \sim P_{\mathcal{X}}$ (Assume $\mathbf{0}$ -mean). The variance of its projection onto (normalized) \mathbf{w} is

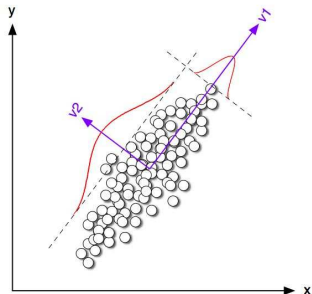
$$\mathbb{E}_{\mathbf{x} \sim \mathcal{X}} \left[\text{proj}_{\mathbf{w}}(\mathbf{x})^2 \right] = \mathbb{E} \left[\mathbf{w}^{\top} \mathbf{x} \mathbf{x}^{\top} \mathbf{w} \right] = \mathbf{w}^{\top} \underbrace{\mathbb{E} \left[\mathbf{x} \mathbf{x}^{\top} \right]}_{\mathbf{C}_{\mathbf{x}\mathbf{x}}} \mathbf{w} = \mathbf{w}^{\top} \mathbf{C}_{\mathbf{x}\mathbf{x}} \mathbf{w}$$

- In input space \mathcal{X} , the **empirical covariance matrix** (of centered data) is

$$\hat{\mathbf{C}}_{\mathbf{x},\mathbf{x}} = \frac{1}{N} \mathbf{X}^{\top} \mathbf{X} ;$$

an $D \times D$ matrix

- How can we find directions that maximize $\mathbf{w}^{\top} \mathbf{C}_{\mathbf{x}\mathbf{x}} \mathbf{w}$? How can we kernelize it?



Recall: Eigenvalues & Eigenvectors



- Given an $N \times N$ matrix \mathbf{A} , an **eigenvector** of \mathbf{A} is a *non-trivial* vector \mathbf{v} that satisfies $\mathbf{A}\mathbf{v} = \lambda\mathbf{v}$; the corresponding value λ is an **eigenvalue**
- Eigen-values/vector pairs satisfy **Rayleigh quotients**:

$$\lambda = \frac{\mathbf{v}^\top \mathbf{A} \mathbf{v}}{\mathbf{v}^\top \mathbf{v}} \qquad \lambda_1 = \max_{\|\mathbf{x}\|=1} \frac{\mathbf{x}^\top \mathbf{A} \mathbf{x}}{\mathbf{x}^\top \mathbf{x}}$$

- Eigen-vectors/values form **orthonormal** matrix \mathbf{V} & diagonal matrix $\mathbf{\Lambda}$

$$\mathbf{V} = \begin{bmatrix} | & | & \dots & | \\ \mathbf{v}_1 & \mathbf{v}_2 & \dots & \mathbf{v}_N \\ | & | & & | \end{bmatrix} \qquad \mathbf{\Lambda} = \begin{bmatrix} \lambda_1(\mathbf{A}) & 0 & \dots & 0 \\ 0 & \lambda_2(\mathbf{A}) & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & & \lambda_N(\mathbf{A}) \end{bmatrix}$$

which form the **eigen-decomposition** of \mathbf{A} : $\mathbf{A} = \mathbf{V}\mathbf{\Lambda}\mathbf{V}^\top$

- Deflation**: for any eigen-value/vector pair (λ, \mathbf{v}) of \mathbf{A} , the transform

$$\tilde{\mathbf{A}} \leftarrow \mathbf{A} - \lambda \mathbf{v} \mathbf{v}^\top$$

deflates the matrix; *i.e.*, \mathbf{v} is an eigenvector of $\tilde{\mathbf{A}}$ but has eigenvalue 0



- Principle Components Analysis (PCA) - algorithm for finding the principle axes of a dataset
- PCA finds subspace spanned by $\{\mathbf{u}_i\}$ that maximizes the data's variance:

$$\mathbf{u}_1 = \underset{\|\mathbf{w}\|=1}{\operatorname{argmax}} \mathbf{w}^\top \mathbf{C}_{xx} \mathbf{w} \qquad \mathbf{C}_{xx} = \frac{1}{N} \mathbf{X}^\top \mathbf{X}$$

- This is achieved by computing \mathbf{C}_{xx} 's eigenvectors
 - 1 Compute the data's mean: $\boldsymbol{\mu} = \frac{1}{N} \sum_{i=1}^N \mathbf{x}_i = \frac{1}{N} \mathbf{X}^\top \mathbf{1}_N$
 - 2 Compute the data's covariance: $\mathbf{C}_{xx} = \frac{1}{N} \sum_{i=1}^N (\mathbf{x}_i - \boldsymbol{\mu})(\mathbf{x}_i - \boldsymbol{\mu})^\top$
 - 3 Find its principle axes: $[\mathbf{U}, \boldsymbol{\Lambda}] = \operatorname{eig}(\mathbf{C}_{xx})$
 - 4 Project data $\{\mathbf{x}_i\}$ onto the first K eigenvectors: $\tilde{\mathbf{x}}_i \leftarrow \mathbf{U}_{1:K}^\top (\mathbf{x}_i - \boldsymbol{\mu})$



- Directions found by PCA are orthonormal: $\mathbf{u}_i^\top \mathbf{u}_j = \delta_{i,j}$
- When projected onto the space spanned by $\{\mathbf{u}_i\}$, resulting data has diagonal covariance matrix
- The eigenvalues λ_i are the amount of variance captured by the direction \mathbf{u}_i
- Variance captured by 1st K directions is $\sum_{i=1}^K \lambda_i (\mathbf{C}_{xx})$
- Using all directions, we can completely reconstruct the data in an alternative basis.
- Directions with low eigenvalues $\lambda_i \ll \lambda_1$ correspond to irrelevant aspects of data... often we use top K directions to re-represent the data.



- **Denoising/Compression:** PCA removes the $(D - K)$ -dimensional subspace with the least information. The PCA transform thus retains the most salient information about the data.
- **Correction:** Reconstruction of data that has been damaged or has missing elements
- **Visualization:** The PCA transform produces a small dimensional projection of data which is convenient for visualizing high dimensional datasets
- **Document Analysis:** PCA can be used to find common themes in a set of documents

Application: Eigenfaces for Face Recognition [1]



Application: Eigenfaces for Face Recognition [1]



Part II

Kernel PCA



- PCA works in the primal space, but not all data structure is well-captured by these linear projections
- How can we kernelize PCA?



- Suppose \mathbf{X} is any $N \times D$ matrix
- The eigen-decomposition of PSD matrices $\mathbf{C}_{xx} = \mathbf{X}^T \mathbf{X}$ & $\mathbf{K} = \mathbf{X} \mathbf{X}^T$ are

$$\mathbf{C}_{xx} = \mathbf{U} \mathbf{\Lambda}_D \mathbf{U}^T \qquad \mathbf{K} = \mathbf{V} \mathbf{\Lambda}_N \mathbf{V}^T$$

where \mathbf{U} & \mathbf{V} are orthogonal and $\mathbf{\Lambda}_D$ & $\mathbf{\Lambda}_N$ have the eigenvalues

- Consider any eigen-pair (λ, \mathbf{v}) of \mathbf{K} . . . then $\mathbf{X}^T \mathbf{v}$ is an eigenvector of \mathbf{C}_{xx} :

$$\mathbf{C}_{xx} \mathbf{X}^T \mathbf{v} = \mathbf{X}^T \mathbf{X} \mathbf{X}^T \mathbf{v} = \mathbf{X}^T \mathbf{K} \mathbf{v} = \lambda \mathbf{X}^T \mathbf{v}$$

and $\|\mathbf{X}^T \mathbf{v}\| = \sqrt{\lambda}$. Thus there is an eigenvector of \mathbf{C}_{xx} such that $\mathbf{u} = \frac{1}{\sqrt{\lambda}} \mathbf{X}^T \mathbf{v}$

- In fact, we have the following correspondences:

$$\mathbf{u} = \lambda^{-1/2} \mathbf{X}^T \mathbf{v} \qquad \mathbf{v} = \lambda^{-1/2} \mathbf{X} \mathbf{u}$$

Singular Value Decomposition II



- Further, let $t = \text{rank}(\mathbf{X}) \leq \min[D, N]$. It can be shown that

$$\text{rank}(\mathbf{C}_{xx}) = \text{rank}(\mathbf{K}) = t$$

- The **singular value decomposition (SVD)** of non-square \mathbf{X} is

$$\mathbf{X} = \mathbf{V}\mathbf{\Sigma}\mathbf{U}^T$$

where \mathbf{U} is $D \times D$ & orthogonal, \mathbf{V} is $N \times N$ & orthogonal, and $\mathbf{\Sigma}$ is $N \times D$ with diagonal given by values $\sigma_i = \sqrt{\lambda_i}$

- The SVD is an analog of eigen-decomposition for non-square matrices.
 - \mathbf{X} is non-singular iff all its singular values are non-zero
 - It yields a spectral decomposition:

$$\mathbf{X} = \sum_i \sigma_i \mathbf{v}_i \mathbf{u}_i^T$$

- Matrix-vector multiply $\mathbf{X}\mathbf{w}$ can be viewed as first projecting \mathbf{w} into eigen-space $\{\mathbf{u}_i\}$ of \mathbf{X} , deforming according to its singular values σ_i and reprojecting into N -space using $\{\mathbf{v}_i\}$



- The SVD decomposition of \mathbf{X} showed a duality in eigenvectors of \mathbf{C}_{xx} and \mathbf{K} that allows us to *kernelize* it
- If \mathbf{u}_j is the j^{th} eigenvector of \mathbf{C}_{xx} , then

$$\mathbf{u}_j = \lambda_j^{-1/2} \mathbf{X}^\top \mathbf{v}_j = \lambda_j^{-1/2} \sum_{i=1}^N \mathbf{X}_{i,\bullet} v_{j,i}$$

i.e., a linear combination of the data points

- Replacing $\mathbf{X}_{i,\bullet}$ with $\phi(\mathbf{x}_i)$, the eigenvector \mathbf{u}_j in feature space is

$$\mathbf{u}_j = \lambda_j^{-1/2} \sum_{i=1}^N v_{j,i} \phi(\mathbf{x}_i) = \sum_{i=1}^N \alpha_{j,i} \phi(\mathbf{x}_i)$$

$$\boldsymbol{\alpha}_j = \lambda_j^{-1/2} \mathbf{v}_j$$

with $\boldsymbol{\alpha}_j$ acting as a *dual vector* defined by eigen-vector \mathbf{v}_j of the *kernel matrix* \mathbf{K}

Projections into Feature Space



- Suppose $\mathbf{u}_j = \sum_{i=1}^N \alpha_{j,i} \phi(\mathbf{x}_i)$ is a normalized direction in the feature space
- For any data point \mathbf{x} , the projection of $\phi(\mathbf{x})$ onto \mathbf{u}_j is

$$\|\text{proj}_{\mathbf{u}_j}(\phi(\mathbf{x}))\| = \mathbf{u}_j^\top \phi(\mathbf{x}) = \sum_{i=1}^N \alpha_{j,i} \kappa(\mathbf{x}_i, \mathbf{x})$$

which represents the *value* of $\phi(\mathbf{x})$ in terms of the j^{th} axis

- Thus, if we have a set of K orthonormal basis vectors $\{\mathbf{u}_j\}_{j=1}^K$, the projection of $\phi(\mathbf{x})$ onto each would produce a new K -vector—

$$\tilde{\mathbf{x}} = \begin{bmatrix} \|\text{proj}_{\mathbf{u}_1}(\phi(\mathbf{x}))\| \\ \|\text{proj}_{\mathbf{u}_2}(\phi(\mathbf{x}))\| \\ \vdots \\ \|\text{proj}_{\mathbf{u}_K}(\phi(\mathbf{x}))\| \end{bmatrix}$$

the representation of $\phi(\mathbf{x})$ in that basis

- Thus, we can perform the PCA transform *in feature space*



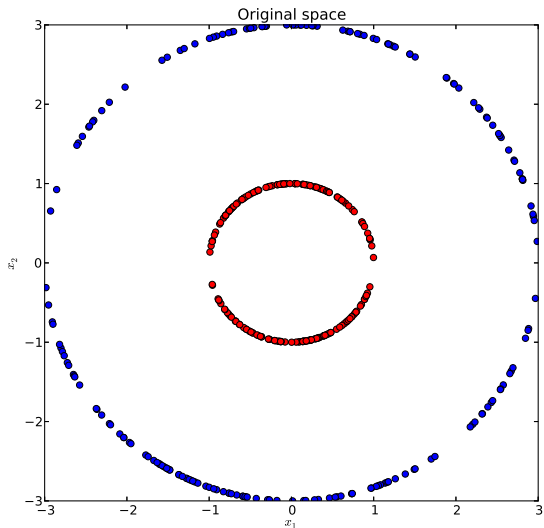
- Performing PCA directly in feature space is not feasible since the covariance matrix is $D \times D$
- However, duality between \mathbf{C}_{xx} & \mathbf{K} allows us to perform PCA indirectly
- Projecting data onto 1st K directions yields a K -dimensional representation
- The algorithm is thus

① Center kernel matrix: $\hat{\mathbf{K}} = \mathbf{K} - \frac{1}{N}\mathbf{1}\mathbf{1}^\top\mathbf{K} - \frac{1}{N}\mathbf{K}\mathbf{1}\mathbf{1}^\top + \frac{\mathbf{1}^\top\mathbf{K}\mathbf{1}}{N^2}\mathbf{1}\mathbf{1}^\top$

② Find its eigenvectors: $[\mathbf{V}, \mathbf{\Lambda}] = \text{eig}(\hat{\mathbf{K}})$

③ Find dual vectors: $\alpha_j = \lambda_j^{-1/2} \mathbf{v}_j$

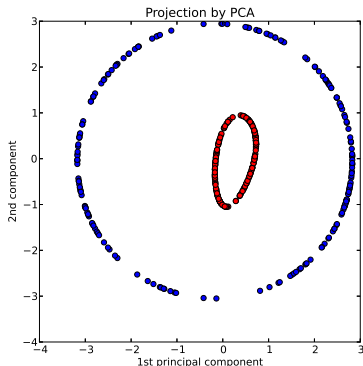
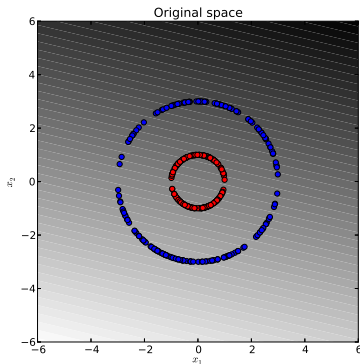
④ Project data onto subspace: $\tilde{\mathbf{x}} \leftarrow \left(\sum_{i=1}^N \alpha_{j,i} \kappa(\mathbf{x}_i, \mathbf{x}) \right)_{j=1}^K$



Kernel PCA - Application



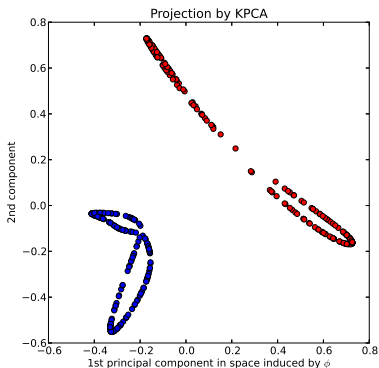
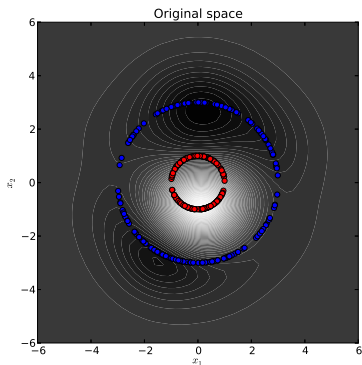
Usual PCA fails to capture the data's two ring structure—the rings are not separated in the first two components.



Kernel PCA - Application



Kernel PCA (RBF) does capture the data's two ring structure & the resulting projections separate the two rings



Part III

Maximum Covariance Analysis



- Suppose we have a pair of related variables: input variable $\mathbf{x} \sim P_{\mathcal{X}}$ and output variable $\mathbf{y} \sim P_{\mathcal{Y}}$ —paired data
- We'd like to find directions of high covariance in spaces $\mathbf{w}_x \in \mathcal{X}$ and $\mathbf{w}_y \in \mathcal{Y}$ such that changes in direction \mathbf{w}_x yield changes in \mathbf{w}_y
- Assuming mean-centered variables, we again have that the covariance of its projection onto (normalized) \mathbf{w}_x & \mathbf{w}_y is

$$\mathbb{E}_{\mathbf{x} \sim \mathcal{X}, \mathbf{y} \sim \mathcal{Y}} \left[\mathbf{w}_x^\top \mathbf{x} \mathbf{w}_y^\top \mathbf{y} \right] = \mathbf{w}_x^\top \underbrace{\mathbb{E} \left[\mathbf{x} \mathbf{y}^\top \right]}_{\mathbf{C}_{xy}} \mathbf{w}_y = \mathbf{w}_x^\top \mathbf{C}_{xy} \mathbf{w}_y$$

- The empirical covariance matrix (of centered data) is

$$\hat{\mathbf{C}}_{\mathbf{x}, \mathbf{y}} = \frac{1}{N} \mathbf{X}^\top \mathbf{Y} ;$$

an $D_{\mathcal{X}} \times D_{\mathcal{Y}}$ matrix

- How can we find directions that maximize $\mathbf{w}_x^\top \mathbf{C}_{xy} \mathbf{w}_y$ for non-square, non-symmetric matrix? How can we kernelize it in space \mathcal{X} ?



- PCA captures structure in data \mathbf{X} , but what data is paired (\mathbf{x}, y) ? We would like to find correlated directions in X and Y
- Suppose we project \mathbf{x} onto direction \mathbf{w}_x and y onto direction \mathbf{w}_y ... the covariance of these random variables is

$$\mathbb{E} \left[\mathbf{w}_x^\top \mathbf{x} \mathbf{w}_y^\top y \right] = \mathbf{w}_x^\top \mathbb{E} \left[\mathbf{x} y^\top \right] \mathbf{w}_y = \mathbf{w}_x^\top \mathbf{C}_{xy} \mathbf{w}_y$$

- The problem we want to solve can again be cast as

$$\max_{\|\mathbf{w}_x\|=1, \|\mathbf{w}_y\|=1} \frac{1}{N} \mathbf{w}_x^\top \mathbf{X}^\top \mathbf{Y} \mathbf{w}_y$$

that is, finding a pair of directions to maximize the covariance

- The solution is simply the first singular vectors $\mathbf{w}_x = \mathbf{u}_1$ & $\mathbf{w}_y = \mathbf{v}_1$ of the SVD $\mathbf{C}_{xy} = \mathbf{U} \mathbf{\Sigma} \mathbf{V}^\top$. Naturally, singular vectors $(\mathbf{u}_2, \mathbf{v}_2), (\mathbf{u}_3, \mathbf{v}_3), \dots$ capture additional covariance



- As with PCA, MCA can also be kernelized by projecting $\mathbf{x} \rightarrow \phi(\mathbf{x})$
- Consider that eigen-analysis of $\mathbf{C}_{xy}\mathbf{C}_{xy}^\top$ gives us \mathbf{U} & of $\mathbf{C}_{xy}^\top\mathbf{C}_{xy}$ gives us \mathbf{V} of the SVD of \mathbf{C}_{xy} ... in fact

$$\mathbf{C}_{xy}^\top\mathbf{C}_{xy} = \frac{1}{N^2}\mathbf{Y}^\top\mathbf{K}_{xx}\mathbf{Y}$$

which has dimension $D_y \times D_y$ & eigen-analysis of this matrix yields (kernelized) directions \mathbf{v}_k

- Then, in decomposing $\mathbf{C}_{xy}\mathbf{C}_{xy}^\top$, we have again a relationship between \mathbf{u}_k & \mathbf{v}_k : $\mathbf{u}_k = \frac{1}{\sigma_k}\mathbf{C}_{xy}\mathbf{v}_k$, allowing us to project onto \mathbf{u}_k when X is kernelized:

$$\|\text{proj}_{\mathbf{u}_k}(\phi(\mathbf{x}))\| = \sum_{i=1}^N \alpha_{k,i} \kappa(\mathbf{x}_i, \mathbf{x}) \quad \alpha_k = \frac{1}{N\sigma_k} \mathbf{Y}\mathbf{v}_k$$

Part IV

Generalized Eigenvalues & CCA

Motivation: Directions of Correlation



- Suppose that instead of input & output variables, we have 2 variables that are different representations of the same data \mathbf{x} :

$$\mathbf{x}_a \leftarrow \psi_a(\mathbf{x}) \qquad \mathbf{x}_b \leftarrow \psi_b(\mathbf{x})$$

- We'd like to find directions of high **correlation** in these spaces $\mathbf{w}_a \in \mathcal{X}_a$ and $\mathbf{w}_b \in \mathcal{X}_b$ such that changes in direction \mathbf{w}_a yield changes in \mathbf{w}_b
- Assuming mean-centered variables, we have that the correlation of its projection onto (normalized) \mathbf{w}_a & \mathbf{w}_b is

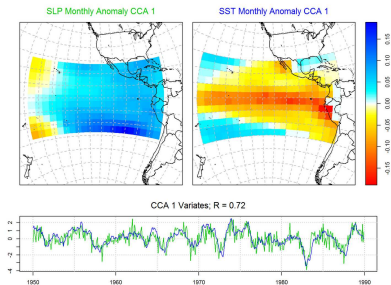
$$\rho_{ab} = \frac{\mathbb{E}_{\mathbf{x}_a \sim \mathcal{X}, \mathbf{x}_b \sim \mathcal{X}_b} [\mathbf{w}_a^\top \mathbf{x}_a \mathbf{w}_b^\top \mathbf{x}_b]}{\sqrt{\mathbb{E} [\mathbf{w}_a^\top \mathbf{x}_a \mathbf{w}_a^\top \mathbf{x}_a] \mathbb{E} [\mathbf{w}_b^\top \mathbf{x}_b \mathbf{w}_b^\top \mathbf{x}_b]}} = \frac{\mathbf{w}_a^\top \mathbf{C}_{ab} \mathbf{w}_b}{\sqrt{\mathbf{w}_a^\top \mathbf{C}_{aa} \mathbf{w}_a \cdot \mathbf{w}_b^\top \mathbf{C}_{bb} \mathbf{w}_b}}$$

where \mathbf{C}_{ab} , \mathbf{C}_{aa} & \mathbf{C}_{bb} are the covariance matrices between \mathbf{x}_a & \mathbf{x}_b (with usual empirical versions)

- How can we find directions that maximize ρ_{ab} ? How can we kernelize it in spaces \mathcal{X}_a & \mathcal{X}_b ?



- Climate Prediction: Researchers have used CCA techniques to find correlations in sea level pressure & sea surface temperature:



- CCA is used with bilingual corpora (same text in two languages) aiding in translation tasks.



- Our objective is to find directions of maximal correlation:

$$\max_{\mathbf{w}_a, \mathbf{w}_b} \rho_{ab}(\mathbf{w}_a, \mathbf{w}_b) = \frac{\mathbf{w}_a^\top \mathbf{C}_{ab} \mathbf{w}_b}{\sqrt{\mathbf{w}_a^\top \mathbf{C}_{aa} \mathbf{w}_a \cdot \mathbf{w}_b^\top \mathbf{C}_{bb} \mathbf{w}_b}} \quad (2)$$

a problem we call **canonical correlation analysis (CCA)**

- As with previous problems this can be expressed as

$$\begin{aligned} \max_{\mathbf{w}_a, \mathbf{w}_b} \quad & \mathbf{w}_a^\top \mathbf{C}_{ab} \mathbf{w}_b \\ \text{such that} \quad & \mathbf{w}_a^\top \mathbf{C}_{aa} \mathbf{w}_a = 1 \text{ and } \mathbf{w}_b^\top \mathbf{C}_{bb} \mathbf{w}_b = 1 \end{aligned} \quad (3)$$



- The Lagrangian function for this optimization is

$$\mathcal{L}(\mathbf{w}_a, \mathbf{w}_b, \lambda_a, \lambda_b) = \mathbf{w}_a^\top \mathbf{C}_{ab} \mathbf{w}_b - \frac{\lambda_a}{2} (\mathbf{w}_a^\top \mathbf{C}_{aa} \mathbf{w}_a - 1) - \frac{\lambda_b}{2} (\mathbf{w}_b^\top \mathbf{C}_{bb} \mathbf{w}_b - 1)$$

- Differentiating it w.r.t. \mathbf{w}_a & \mathbf{w}_b & setting equal to 0 gives

$$\mathbf{C}_{ab} \mathbf{w}_b - \lambda_a \mathbf{C}_{aa} \mathbf{w}_a = 0 \qquad \mathbf{C}_{ba} \mathbf{w}_a - \lambda_b \mathbf{C}_{bb} \mathbf{w}_b = 0$$

$$\lambda_a \mathbf{w}_a^\top \mathbf{C}_{aa} \mathbf{w}_a = \lambda_b \mathbf{w}_b^\top \mathbf{C}_{bb} \mathbf{w}_b$$

which implies that $\lambda_a = \lambda_b = \lambda$

- The constraints on \mathbf{w}_a & \mathbf{w}_b can be written in matrix form as

$$\begin{bmatrix} \mathbf{0} & \mathbf{C}_{ab} \\ \mathbf{C}_{ba} & \mathbf{0} \end{bmatrix} \begin{bmatrix} \mathbf{w}_a \\ \mathbf{w}_b \end{bmatrix} = \lambda \begin{bmatrix} \mathbf{C}_{aa} & \mathbf{0} \\ \mathbf{0} & \mathbf{C}_{bb} \end{bmatrix} \begin{bmatrix} \mathbf{w}_a \\ \mathbf{w}_b \end{bmatrix} \quad (4)$$

$$\mathbf{A} \mathbf{w} = \lambda \mathbf{B} \mathbf{w} ;$$

a **generalized eigenvalue problem** for the primal problem



- Suppose \mathbf{A} & \mathbf{B} are symmetric & $\mathbf{B} \succ 0$, then the **generalized eigenvalue problem (GEP)** is to find (λ, \mathbf{w}) s.t.

$$\mathbf{A}\mathbf{w} = \lambda\mathbf{B}\mathbf{w} \quad (5)$$

which are equivalent to

$$\max_{\mathbf{w}} \frac{\mathbf{w}^{\top} \mathbf{A} \mathbf{w}}{\mathbf{w}^{\top} \mathbf{B} \mathbf{w}} \qquad \max_{\mathbf{w}^{\top} \mathbf{B} \mathbf{w} = 1} \mathbf{w}^{\top} \mathbf{A} \mathbf{w}$$

Note, eigenvalues are special case with $\mathbf{B} = \mathbf{I}$

- Since $\mathbf{B} \succ 0$, any GEP can be converted to an Eigenvalue problem by inverting \mathbf{B} :

$$\mathbf{B}^{-1} \mathbf{A} \mathbf{w} = \lambda \mathbf{w}$$



- However, to ensure symmetry, we can instead use $\mathbf{B} \succ 0$ to decompose $\mathbf{B} = \mathbf{B}^{-1/2} \mathbf{B}^{-1/2}$ where $\mathbf{B}^{-1/2} = \sqrt{\mathbf{B}^{-1}}$ is a symmetric real matrix—taking $\mathbf{w} = \mathbf{B}^{-1/2} \mathbf{v}$ for some \mathbf{v} we obtain (symmetric)

$$\mathbf{B}^{-1/2} \mathbf{A} \mathbf{B}^{-1/2} \mathbf{v} = \lambda \mathbf{v}$$

an eigenvalue problem for $\mathbf{C} = \mathbf{B}^{-1/2} \mathbf{A} \mathbf{B}^{-1/2}$ providing solutions to Eq. (5)

$$\mathbf{w}_i = \mathbf{B}^{-1/2} \mathbf{v}_i$$



Proposition 1

Solutions to GEP of Eq. (5) have following properties: if eigenvalues are distinct, then

$$\mathbf{w}_i^\top \mathbf{B} \mathbf{w}_j = \delta_{i,j}$$

$$\mathbf{w}_i^\top \mathbf{A} \mathbf{w}_j = \lambda_i \delta_{i,j}$$

*that is, the vectors \mathbf{w}_i are orthonormal after applying transformation $\mathbf{B}^{1/2}$ —that is, they are **conjugate** with respect to \mathbf{B} .*



Theorem 2

If $(\lambda_i, \mathbf{w}_i)$ are eigen-solutions to GEP of Eq. (5), then \mathbf{A} can be decomposed as

$$\mathbf{A} = \sum_{i=1}^N \lambda_i \mathbf{B} \mathbf{w}_i (\mathbf{B} \mathbf{w}_i)^\top$$

This yields the *generalized deflation* of \mathbf{A} :

$$\tilde{\mathbf{A}} \leftarrow \mathbf{A} - \lambda_i \mathbf{B} \mathbf{w}_i \mathbf{w}_i^\top \mathbf{B}^\top$$

while \mathbf{B} is unchanged.



- As shown in Eq. (4), CCA is a GEP $\mathbf{A}\mathbf{w} = \lambda\mathbf{B}\mathbf{w}$ where

$$\mathbf{A} = \begin{bmatrix} \mathbf{0} & \mathbf{C}_{ab} \\ \mathbf{C}_{ba} & \mathbf{0} \end{bmatrix} \quad \mathbf{B} = \begin{bmatrix} \mathbf{C}_{aa} & \mathbf{0} \\ \mathbf{0} & \mathbf{C}_{bb} \end{bmatrix} \quad \mathbf{w} = \begin{bmatrix} \mathbf{w}_a \\ \mathbf{w}_b \end{bmatrix}$$

- Since this is a solution to Eq. (2), the eigenvalues will be **correlations** $\Rightarrow \lambda \in [-1, +1]$. Further, the eigensolutions will pair: for each $\lambda_i > 0$ with eigenvector $\begin{bmatrix} \mathbf{w}_a \\ \mathbf{w}_b \end{bmatrix}$, there is a $\lambda_j = -\lambda_i$ with eigenvector $\begin{bmatrix} \mathbf{w}_a \\ -\mathbf{w}_b \end{bmatrix}$. *Hence, we only need to consider the positive spectrum.*
- Larger eigenvalues correspond to the strongest correlations.
- Finally, the solutions are **conjugate** w.r.t. matrix \mathbf{B} which reveals that for $i \neq j$

$$\mathbf{w}_{a,j}^\top \mathbf{C}_{aa} \mathbf{w}_{a,i} = 0$$

$$\mathbf{w}_{b,j}^\top \mathbf{C}_{bb} \mathbf{w}_{b,i} = 0$$

However, the directions will not be orthogonal in the original input space.



- Let's take the directions to be linear combinations of data:

$$\mathbf{w}_a = \mathbf{X}_a^\top \boldsymbol{\alpha}_a \qquad \mathbf{w}_b = \mathbf{X}_b^\top \boldsymbol{\alpha}_b$$

- Substituting these directions into Eq. (3) gives

$$\begin{aligned} & \max_{\boldsymbol{\alpha}_a, \boldsymbol{\alpha}_b} \boldsymbol{\alpha}_a^\top \mathbf{K}_a \mathbf{K}_b \boldsymbol{\alpha}_b \\ & \text{such that } \boldsymbol{\alpha}_a^\top \mathbf{K}_a^2 \boldsymbol{\alpha}_a = 1 \text{ and } \boldsymbol{\alpha}_b^\top \mathbf{K}_b^2 \boldsymbol{\alpha}_b = 1 \end{aligned}$$

where $\mathbf{K}_a = \mathbf{X}_a \mathbf{X}_a^\top$ and $\mathbf{K}_b = \mathbf{X}_b \mathbf{X}_b^\top$.



- Differentiating the Lagrangian again yields equations

$$\mathbf{K}_a \mathbf{K}_b \alpha_b - \lambda \mathbf{K}_a^2 \alpha_a = \mathbf{0} \qquad \mathbf{K}_b \mathbf{K}_a \alpha_a - \lambda \mathbf{K}_b^2 \alpha_b = \mathbf{0}$$

- However, these equations reveal a problem. When the dimension of the feature space is large compared number of data points ($D_a \gg N$), solutions will overfit the data.
- For the Gaussian kernel, data will always be independent in feature space & \mathbf{K}_a will be invertible. Hence, we have

$$\alpha_a = \frac{1}{\lambda} \mathbf{K}_a^{-1} \mathbf{K}_b \alpha_b$$
$$\mathbf{K}_b^2 \alpha_b - \lambda^2 \mathbf{K}_b^2 \alpha_b = \mathbf{0}$$

but the latter holds for all α_b with perfect correlation $\lambda = 1$ —*Solution is Overfit!!!*



- To avoid overfitting, we can regularize the solutions \mathbf{w}_a & \mathbf{w}_b by controlling their norms. The **Regularized CCA Problem** is

$$\max_{\mathbf{w}_a, \mathbf{w}_b} \tilde{\rho}_{ab}(\mathbf{w}_a, \mathbf{w}_b) = \frac{\mathbf{w}_a^\top \mathbf{C}_{ab} \mathbf{w}_b}{\sqrt{\left((1 - \tau_a) \mathbf{w}_a^\top \mathbf{C}_{aa} \mathbf{w}_a + \tau_a \|\mathbf{w}_a\|^2 \right) \cdot \left((1 - \tau_b) \mathbf{w}_b^\top \mathbf{C}_{bb} \mathbf{w}_b + \tau_b \|\mathbf{w}_b\|^2 \right)}}$$

where $\tau_a \in [0, 1]$ & $\tau_b \in [0, 1]$ serve as regularization parameters

- Again this yields an optimization program for the dual variables

$$\begin{aligned} \max_{\mathbf{w}_a, \mathbf{w}_b} \quad & \alpha_a^\top \mathbf{K}_a \mathbf{K}_b \alpha_b \\ \text{such that} \quad & (1 - \tau_a) \alpha_a^\top \mathbf{K}_a^2 \alpha_a + \tau_a \alpha_a^\top \mathbf{K}_a \alpha_a = 1 \\ & \text{and} \quad (1 - \tau_b) \alpha_b^\top \mathbf{K}_b^2 \alpha_b + \tau_b \alpha_b^\top \mathbf{K}_b \alpha_b = 1 \end{aligned}$$



- Using the Lagrangian technique, we again arrive at a GEP:

$$\begin{bmatrix} \mathbf{0} & \mathbf{K}_a \mathbf{K}_b \\ \mathbf{K}_b \mathbf{K}_a & \mathbf{0} \end{bmatrix} \begin{bmatrix} \alpha_a \\ \alpha_b \end{bmatrix} = \lambda \begin{bmatrix} (1 - \tau_a) \mathbf{K}_a^2 + \tau_a \mathbf{K}_a & \mathbf{0} \\ \mathbf{0} & (1 - \tau_b) \mathbf{K}_b^2 + \tau_b \mathbf{K}_b \end{bmatrix} \begin{bmatrix} \alpha_a \\ \alpha_b \end{bmatrix}$$

- Solutions (α_a^*, α_b^*) can now be used as usual projection directions of Eq. (1)
- Solving CCA using the above GEP is *impractical!* The matrices required are $2N \times 2N$. Instead, the usual approach is to make an **incomplete Cholesky decomposition** of the kernel matrices:

$$\mathbf{K}_a = \mathbf{R}_a^\top \mathbf{R}_a \qquad \mathbf{K}_b = \mathbf{R}_b^\top \mathbf{R}_b$$

The resulting GEP can be solved more efficiently (see book for algorithms details)



- Finally CCA can be extended to multiple representations of the data, which result in the following GEP:

$$\begin{bmatrix} \mathbf{C}_{11} & \mathbf{C}_{12} & \cdots & \mathbf{C}_{1k} \\ \mathbf{C}_{21} & \mathbf{C}_{22} & \cdots & \mathbf{C}_{2k} \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{C}_{k1} & \mathbf{C}_{k2} & \cdots & \mathbf{C}_{kk} \end{bmatrix} \begin{bmatrix} \mathbf{w}_1 \\ \mathbf{w}_2 \\ \vdots \\ \mathbf{w}_k \end{bmatrix} = \rho \begin{bmatrix} \mathbf{C}_{11} & \mathbf{0} & \cdots & \mathbf{0} \\ \mathbf{0} & \mathbf{C}_{22} & \cdots & \mathbf{0} \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{0} & \mathbf{0} & \cdots & \mathbf{C}_{kk} \end{bmatrix} \begin{bmatrix} \mathbf{w}_1 \\ \mathbf{w}_2 \\ \vdots \\ \mathbf{w}_k \end{bmatrix}$$



You should note, that the Fisher Discriminant Analysis problem can be expressed as

$$\max_{\alpha} J(\alpha) = \frac{\alpha^{\top} \mathbf{M} \alpha}{\alpha^{\top} \mathbf{N} \alpha}$$

which is a GEP. In fact, this is how solutions to LDA are obtained.



- In this lecture, we saw how different objectives for projection directions yield different subspaces. . . we saw 3 different algorithms:
 - 1 Principal Component Analysis
 - 2 Maximum Covariance Analysis
 - 3 Canonical Correlation Analysis
- We saw that each of these techniques can be solved using eigenvalue, singular value, and generalized eigenvector decompositions.
- We saw that each of these techniques yielded linear projections and thus could be kernelized.
- In the next lecture, we will explore the general technique of minimizing loss & how allows us to develop a wide range of kernel algorithms. In particular, we will see the **Support Vector Machine** for classification tasks.



The Majority of the work from this talk can be found in the lecture's accompanying book, "Kernel Methods for Pattern Analysis."

- [1] M. A. Turk and A. P. Pentland. Face recognition using eigenfaces. In *IEEE Computer Society Conference on Computer Vision and Pattern Recognition*, pages 586–591, 1991.